

## Today's Discussion

### Review from last time (Sect. 2.1, 2.4)

- Bayesian decision theory
- Discriminant functions

### Normal density: (Sect. 2.5, 2.6)

- Normal density function
- Discriminant functions for normal densities

### Maximum likelihood estimation: (Sect. 3.1, 3.2)

## Review from last time

- Assume knowledge of  $P(\omega_i)$  and  $p(\mathbf{x}|\omega_i)$  for all classes.
- *Bayes Optimal Classifier:*

**Decide**  $\omega_i$  **if**  $P(\omega_i|\mathbf{x}) > P(\omega_j|\mathbf{x})$ ,  $\forall j \neq i$  **where,**

$$P(\omega_i|\mathbf{x}) = \frac{p(\mathbf{x}|\omega_i)P(\omega_i)}{p(\mathbf{x})}$$

$$p(\mathbf{x}) = \sum_{j=1}^k p(\mathbf{x}|\omega_j)P(\omega_j)$$

## Discriminant functions

### Three choices:

$$\text{\#1: } g_i(\mathbf{x}) = P(\omega_i|\mathbf{x}) = \frac{p(\mathbf{x}|\omega_i)P(\omega_i)}{p(\mathbf{x})}$$

$$\text{\#2: } g_i(\mathbf{x}) = p(\mathbf{x}|\omega_i)P(\omega_i)$$

$$\text{\#3: } g_i(\mathbf{x}) = \ln p(\mathbf{x}|\omega_i) + \ln P(\omega_i)$$

### Bayes Decision Rule:

**Decide**  $\omega_i$  **if**  $g_i(\mathbf{x}) > g_j(\mathbf{x})$ ,  $\forall j \neq i$

## Two-category case

### Let:

$$g(\mathbf{x}) \equiv g_1(\mathbf{x}) - g_2(\mathbf{x})$$

### Bayes Optimal Classifier (*Dichotomizer*):

**Decide**  $\omega_1$  **if**  $g(\mathbf{x}) > 0$  ; **else decide**  $\omega_2$  .

$$g(\mathbf{x}) = P(\omega_1|\mathbf{x}) - P(\omega_2|\mathbf{x})$$

$$g(\mathbf{x}) = \ln \frac{p(\mathbf{x}|\omega_1)}{p(\mathbf{x}|\omega_2)} + \ln \frac{P(\omega_1)}{P(\omega_2)}$$

## Normal (Gaussian) density

One-dimensional (univariate) density:

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right]$$

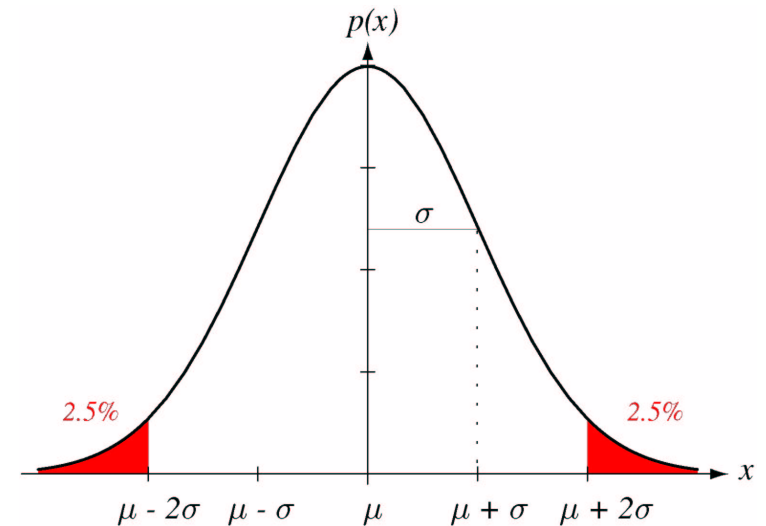
$$\mu \equiv E[x] = \int_{-\infty}^{\infty} xp(x)dx = \text{mean}$$

$$\sigma^2 = E[(x-\mu)^2] = \int_{-\infty}^{\infty} (x-\mu)^2 p(x)dx = \text{variance}$$

$\sigma$  = standard deviation

Let's confirm integrals in *Mathematica* ...

## Univariate density



## Multivariate normal density

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2}(\mathbf{x}-\mu)^T \Sigma^{-1} (\mathbf{x}-\mu)\right]$$

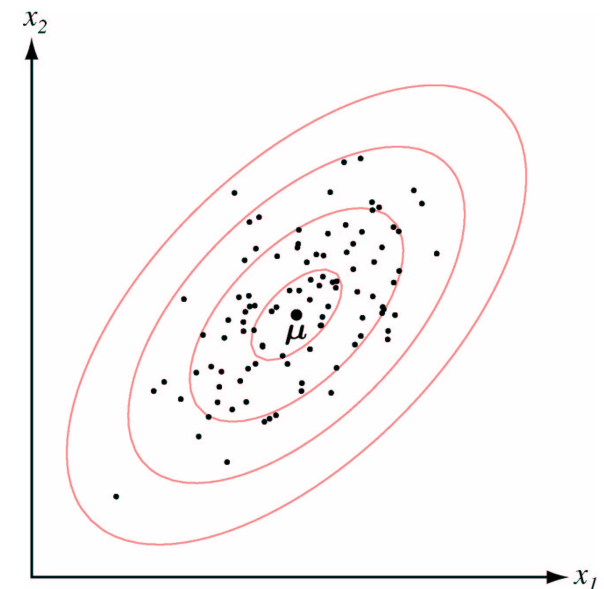
$$\mu \equiv E[\mathbf{x}] = \int \mathbf{x} p(\mathbf{x}) d\mathbf{x} = \text{mean vector}$$

$$\Sigma \equiv E[(\mathbf{x}-\mu)(\mathbf{x}-\mu)^T] = \int (\mathbf{x}-\mu)(\mathbf{x}-\mu)^T p(\mathbf{x}) d\mathbf{x} = \text{covariance matrix (symmetric, } \Sigma \geq 0)$$

$d$  = dimension

$p(\mathbf{x}) \sim N(\mu, \Sigma)$  (alternate notation)

## Multivariate density example



## Discriminant functions for normal density

$$g_i(\mathbf{x}) = \ln p(\mathbf{x}|\omega_i) + \ln P(\omega_i)$$

$$p(\mathbf{x}|\omega_i) = \frac{1}{(2\pi)^{d/2} |\Sigma_i|^{1/2}} \exp\left[-\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i)\right]$$

$$\ln p(\mathbf{x}|\omega_i) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i|$$

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

## Case 1: $\Sigma_i = \sigma^2 \mathbf{I}$

Equal, uniform covariance matrices:

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

$$|\Sigma_i| = \sigma^{2d}, \Sigma_i^{-1} = \frac{1}{\sigma^2} \mathbf{I}$$

$$g_i(\mathbf{x}) = \frac{-1}{2\sigma^2}(\mathbf{x} - \mu_i)^T (\mathbf{x} - \mu_i) - \frac{d}{2} \ln 2\pi - d \ln \sigma + \ln P(\omega_i)$$

$$g_i(\mathbf{x}) = \frac{-1}{2\sigma^2}(\mathbf{x} - \mu_i)^T (\mathbf{x} - \mu_i) + \ln P(\omega_i)$$

## Discriminant function (Case 1)

$$g_i(\mathbf{x}) = \frac{-1}{2\sigma^2}(\mathbf{x} - \mu_i)^T (\mathbf{x} - \mu_i) + \ln P(\omega_i)$$

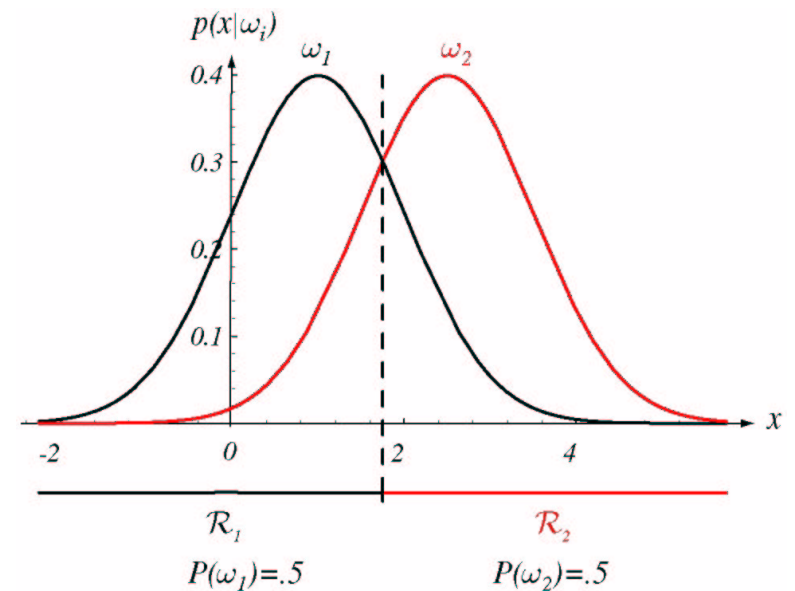
$$g_i(\mathbf{x}) = \frac{-1}{2\sigma^2}[\mathbf{x}^T \mathbf{x} - 2\mu_i^T \mathbf{x} + \mu_i^T \mu_i] + \ln P(\omega_i)$$

$$g_i(\mathbf{x}) = \frac{\mu_i^T}{\sigma^2} \mathbf{x} - \frac{\mu_i^T \mu_i}{2\sigma^2} + \ln P(\omega_i)$$

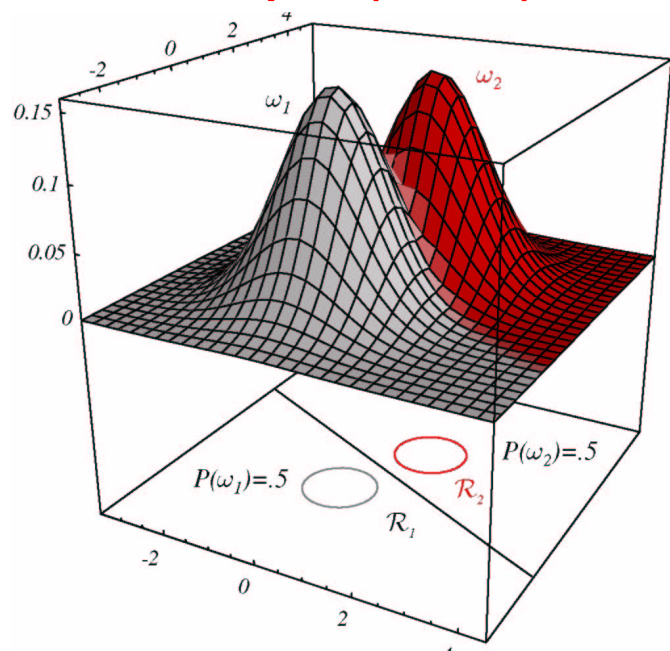
$$g_i(\mathbf{x}) = \mathbf{w}_i^T \mathbf{x} + w_{i0} \Rightarrow \text{linear discriminant function}$$

$$\mathbf{w}_i = \frac{\mu_i}{\sigma^2}, w_{i0} = -\frac{\mu_i^T \mu_i}{2\sigma^2} + \ln P(\omega_i)$$

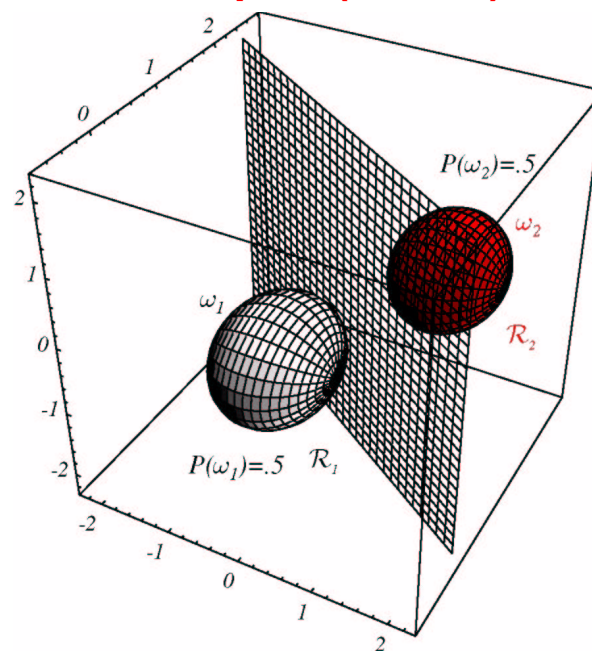
## Examples (Case 1)



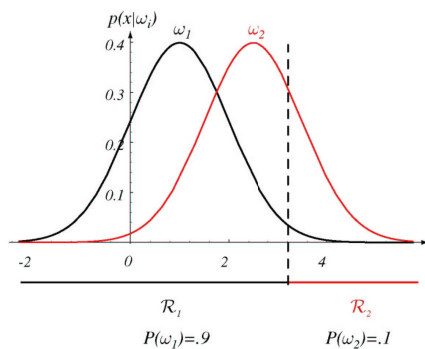
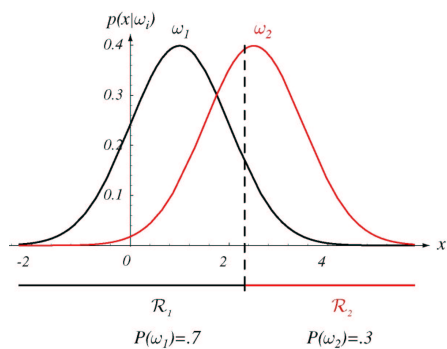
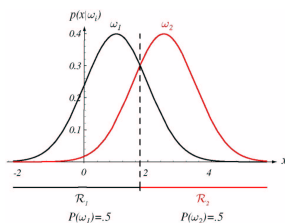
### Examples (Case 1)



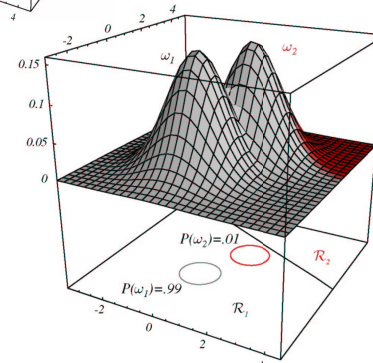
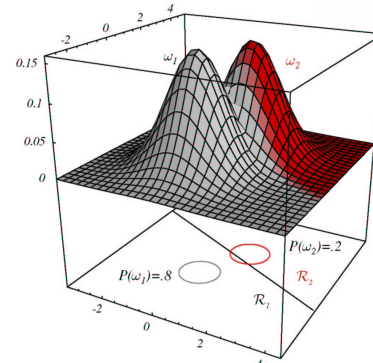
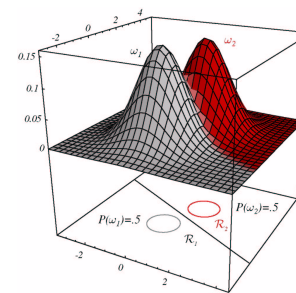
### Examples (Case 1)



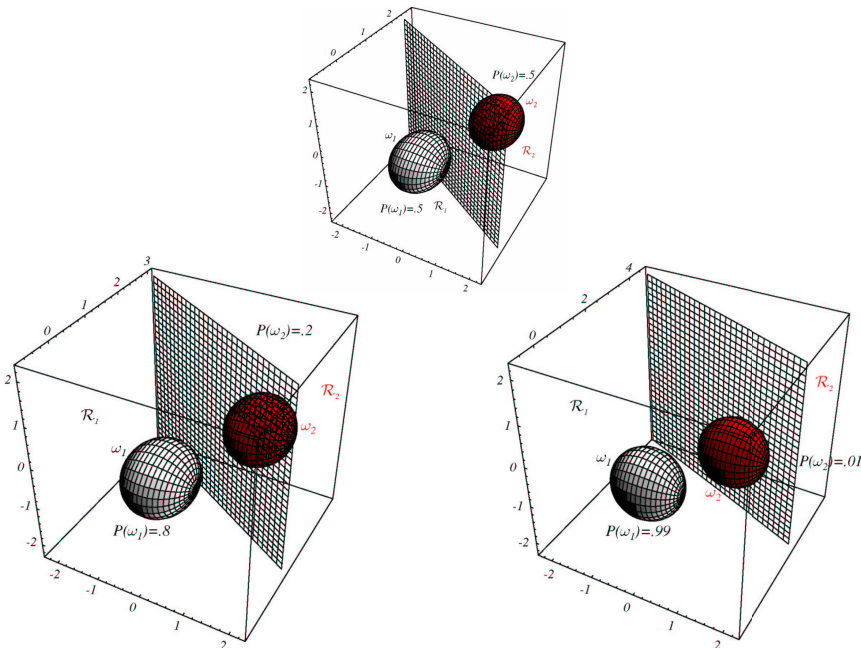
### Non-equal priors (Case 1)



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## Case 2: $\Sigma_i = \Sigma$

Equal covariance matrices:

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma^{-1} (\mathbf{x} - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma| + \ln P(\omega_i)$$

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma^{-1} (\mathbf{x} - \mu_i) + \ln P(\omega_i)$$

## Discriminant function (Case 2)

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma^{-1} (\mathbf{x} - \mu_i) + \ln P(\omega_i)$$

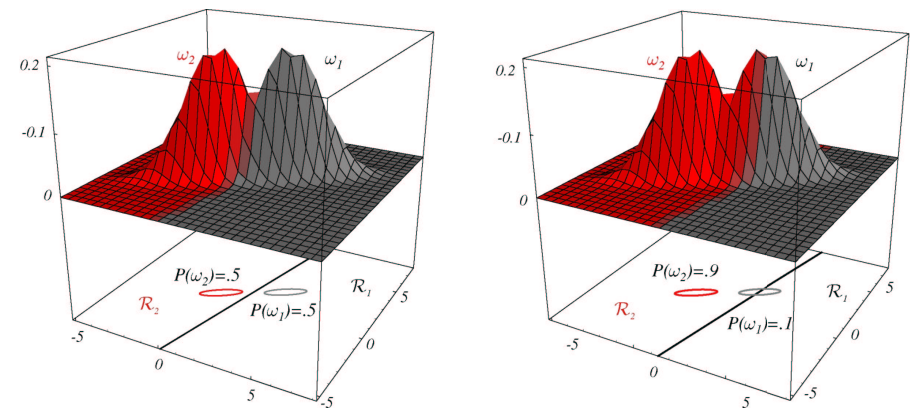
$$g_i(\mathbf{x}) = -\frac{1}{2}[\mathbf{x}^T \Sigma^{-1} \mathbf{x} - 2\mu_i^T \Sigma^{-1} \mathbf{x} + \mu_i^T \Sigma^{-1} \mu_i] + \ln P(\omega_i)$$

$$g_i(\mathbf{x}) = \mu_i^T \Sigma^{-1} \mathbf{x} - \frac{1}{2} \mu_i^T \Sigma^{-1} \mu_i + \ln P(\omega_i)$$

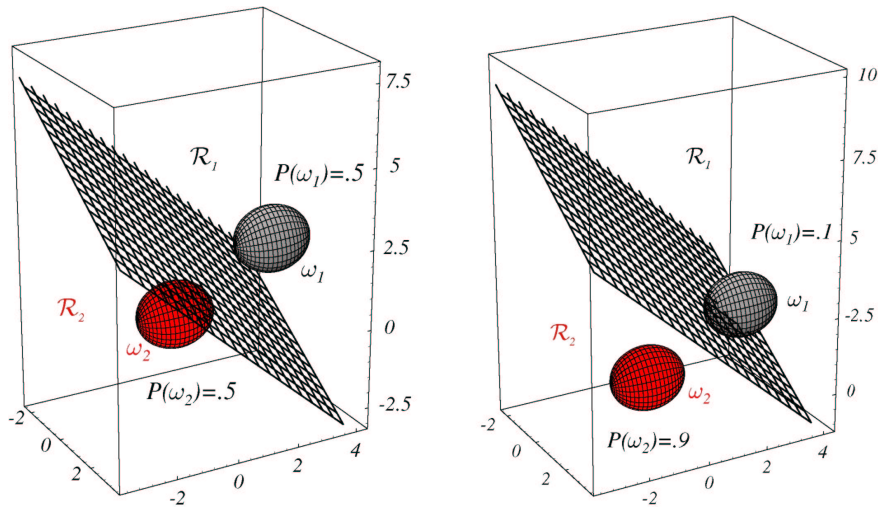
$$g_i(\mathbf{x}) = \mathbf{w}_i^T \mathbf{x} + w_{i0} \Rightarrow \text{linear discriminant function}$$

$$\mathbf{w}_i = \Sigma^{-1} \mu_i, w_{i0} = -\frac{1}{2} \mu_i^T \Sigma^{-1} \mu_i + \ln P(\omega_i)$$

## Examples (Case 2)



## Examples (Case 2)



## Case 3: $\Sigma_i = \text{arbitrary}$

Arbitrary covariance matrices:

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i) - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

## Discriminant function (Case 3)

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i) - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

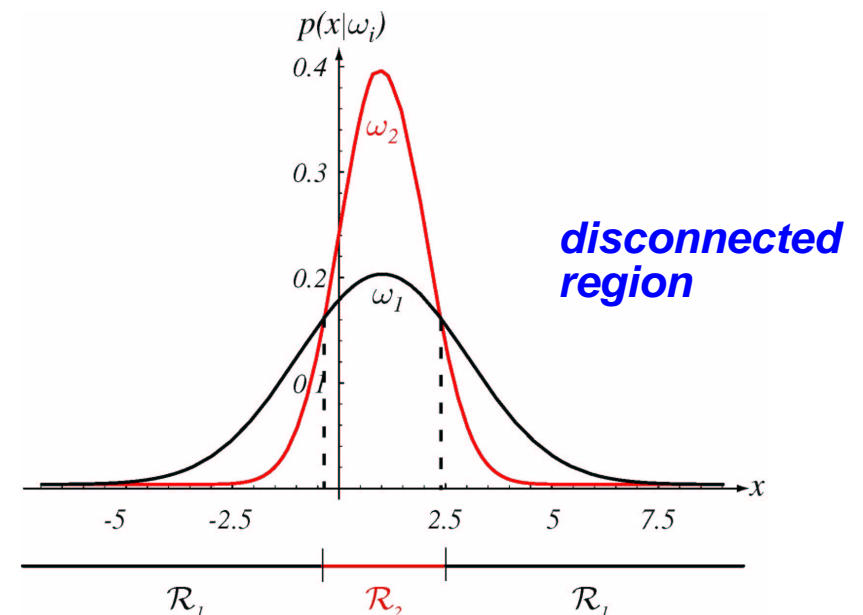
Quadratic discriminant function:

$$g_i(\mathbf{x}) = \mathbf{x}^T \mathbf{W}_i \mathbf{x} + \mathbf{w}_i^T \mathbf{x} + w_{i0}$$

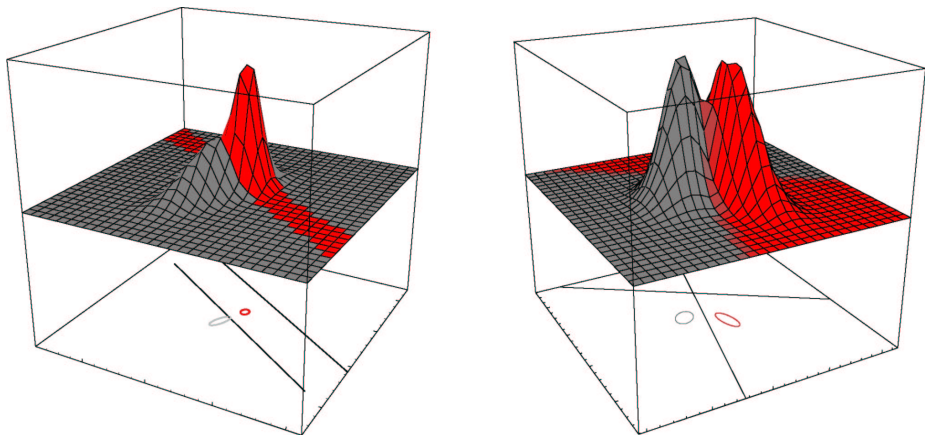
$$\mathbf{W}_i = -\frac{1}{2} \Sigma_i^{-1}, \quad \mathbf{w}_i = \Sigma_i^{-1} \mu_i$$

$$w_{i0} = -\frac{1}{2} \mu_i^T \Sigma_i^{-1} \mu_i - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

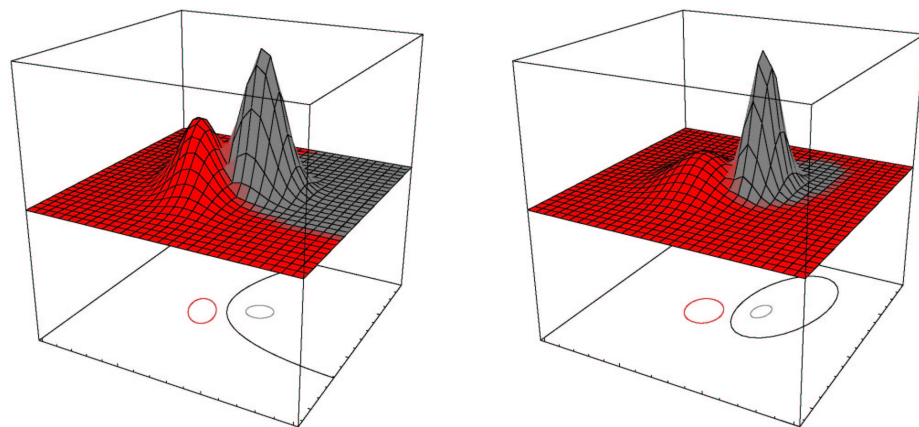
## Case 3 examples



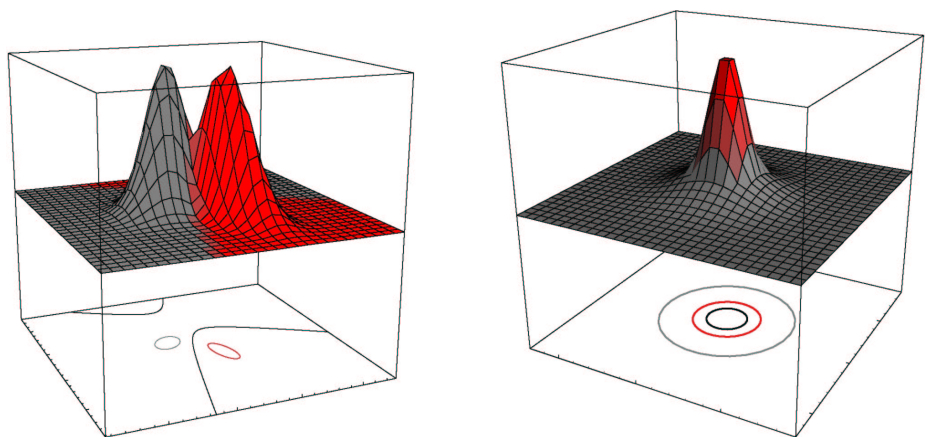
Case 3 examples (2d)



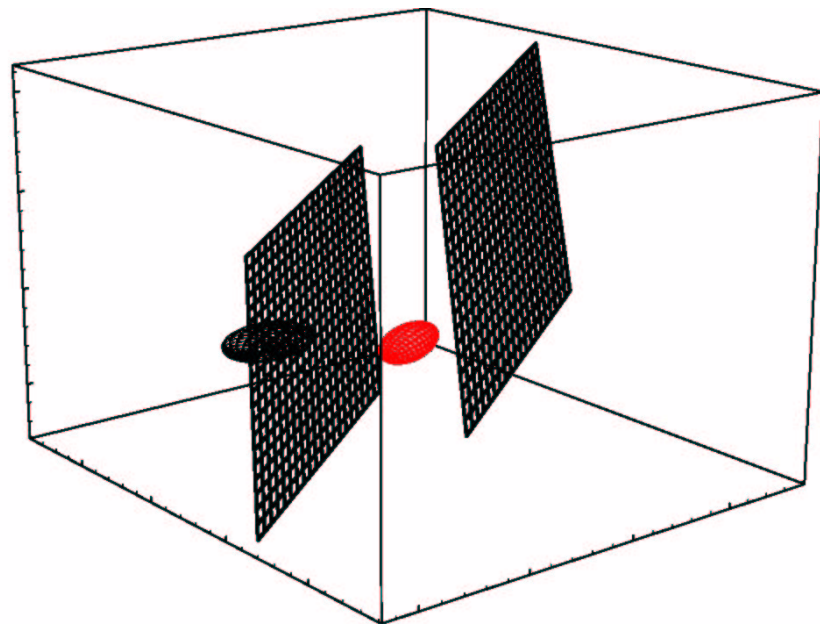
Case 3 examples (2d)



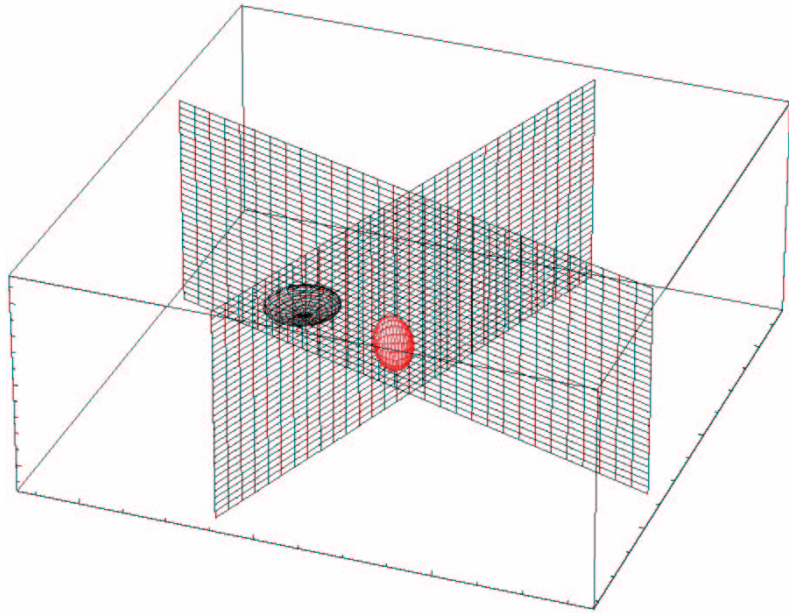
Case 3 examples (2d)



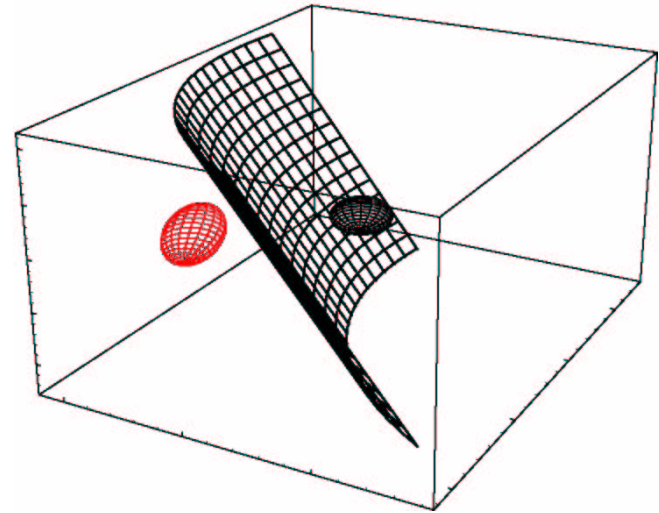
Case 3 examples (3d)



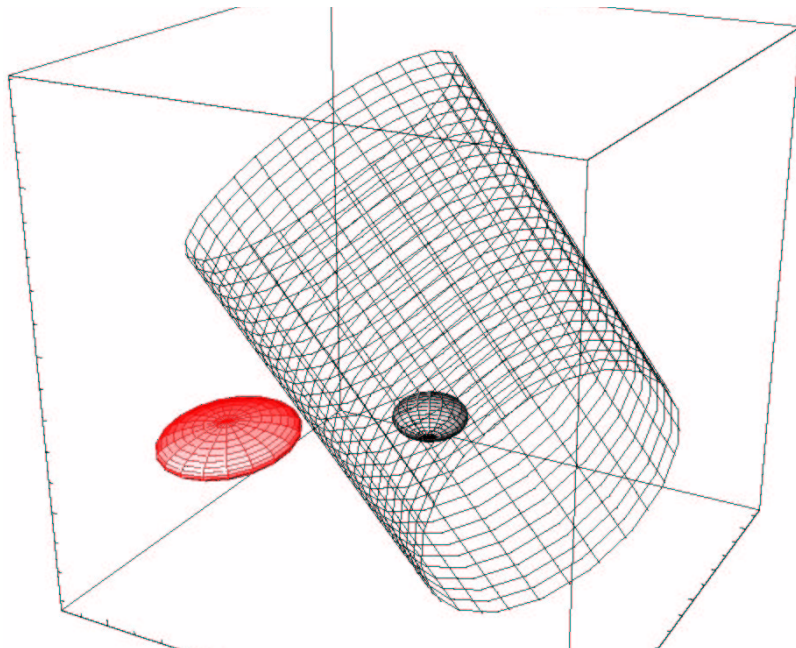
**Case 3 examples (3d)**



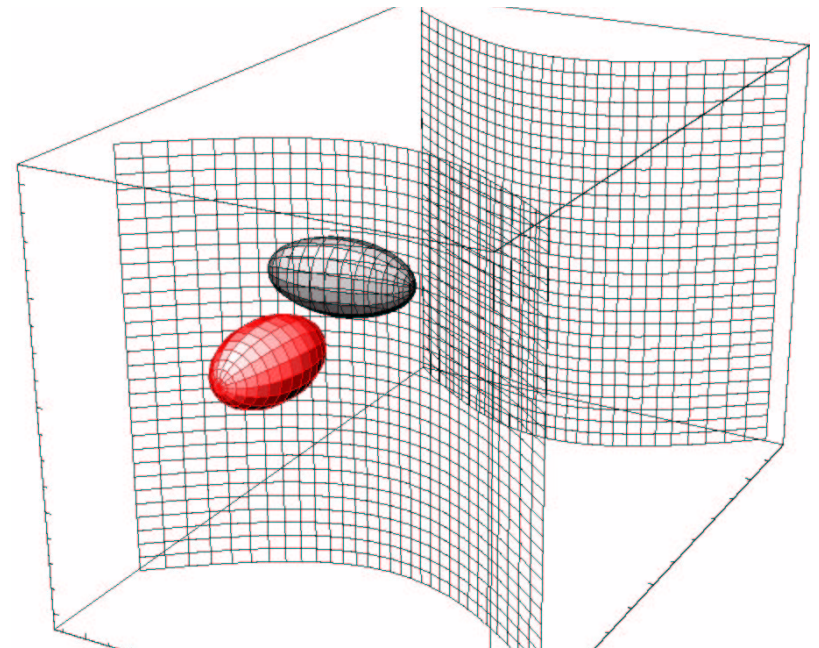
**Case 3 examples (3d)**



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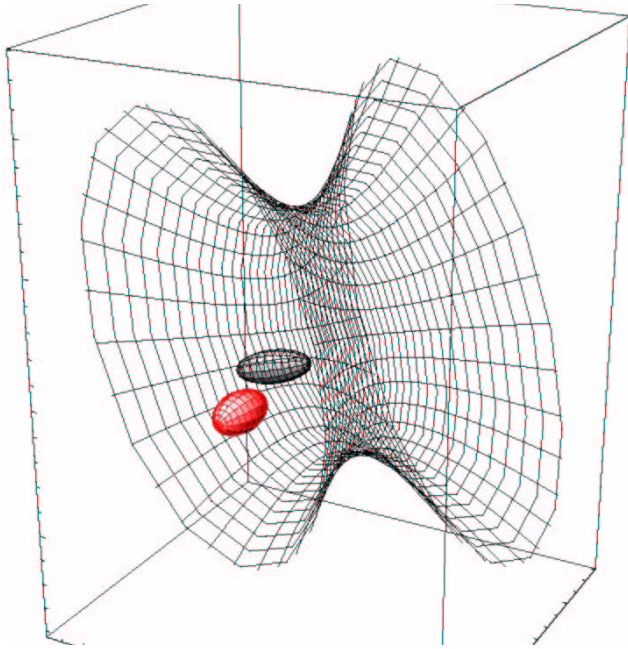


**Case 3 examples (3d)**

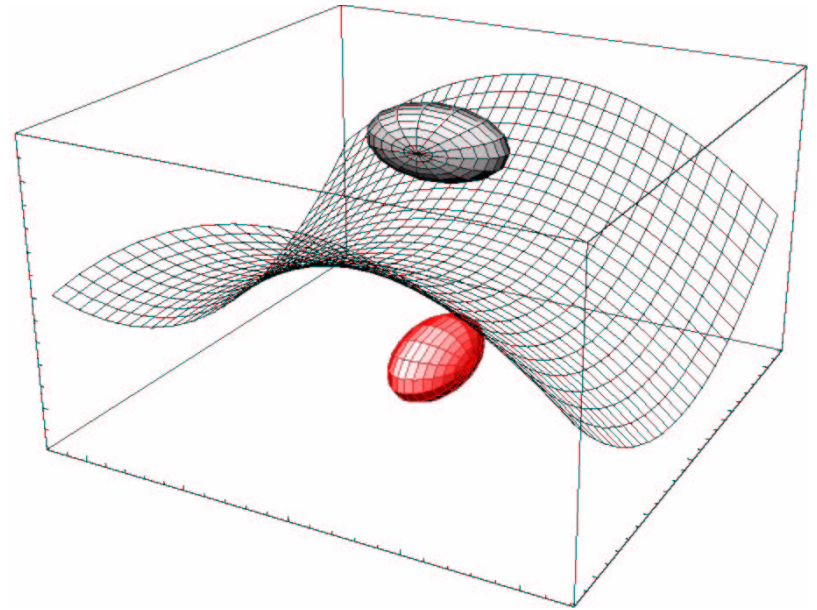




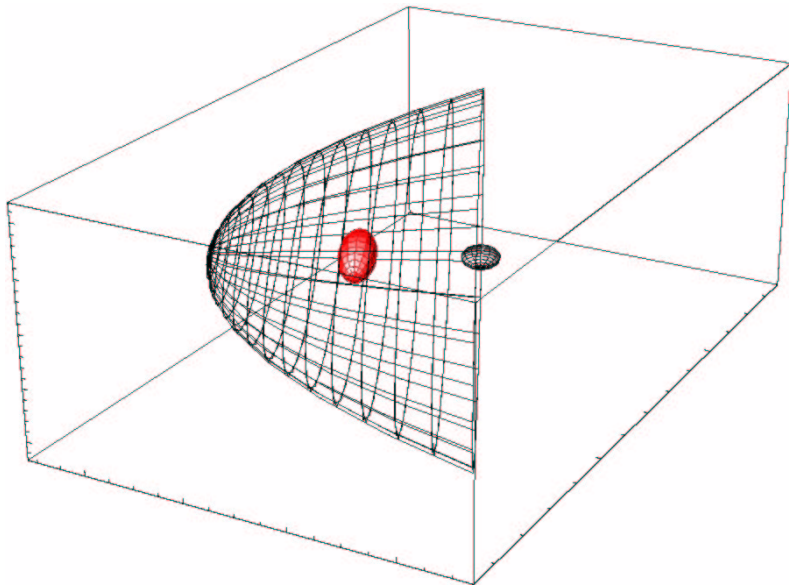
**Case 3 examples (3d)**



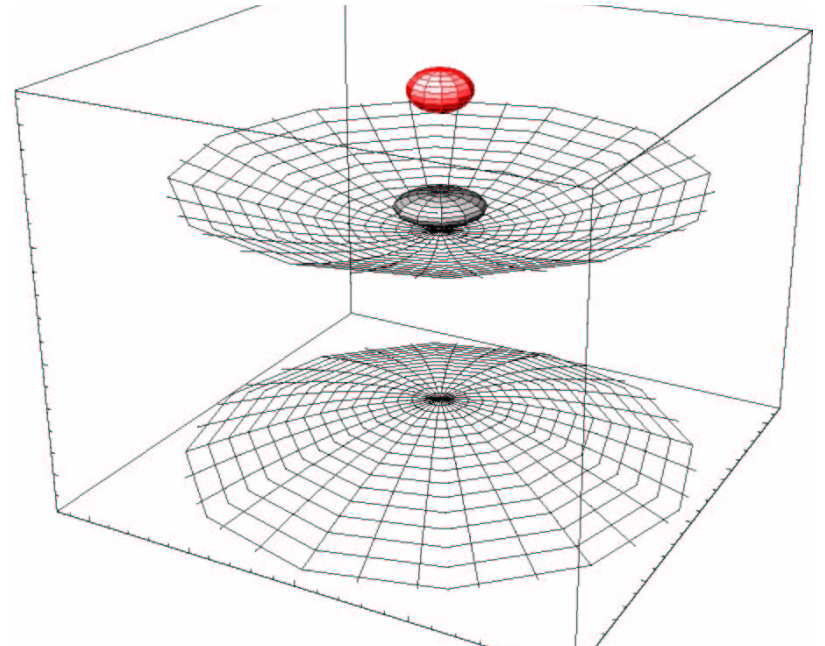
**Case 3 examples (3d)**



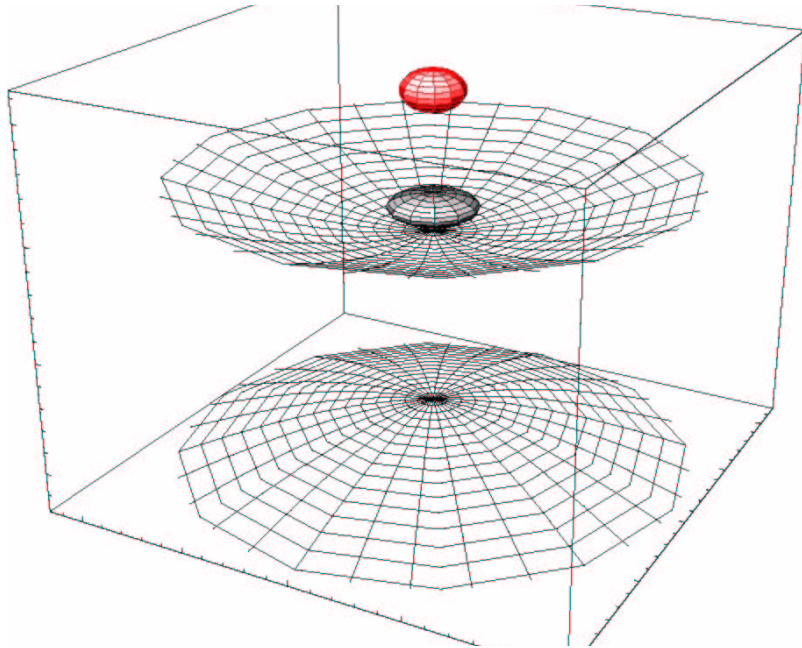
**Case 3 examples (3d)**



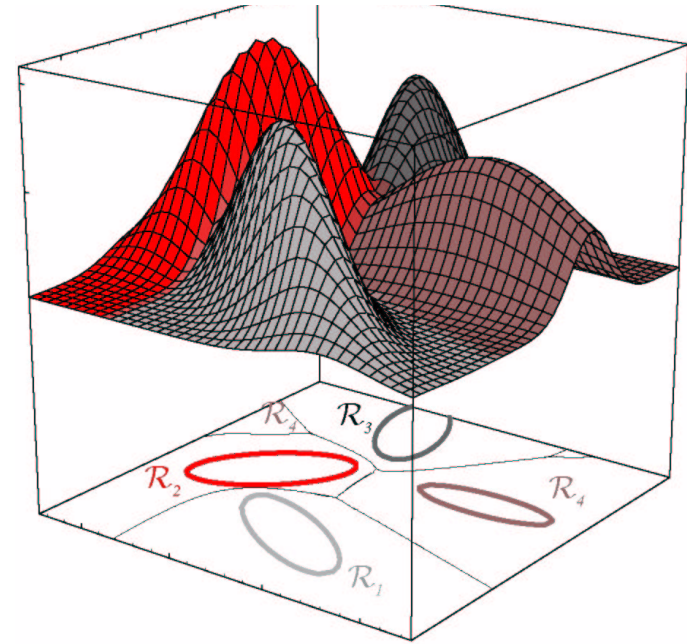
**Case 3 examples (3d)**



### Case 3 examples (3d)



### Multiple regions



### Example of decision boundary derivation

Two normal-distributed classes:  $\omega_1, \omega_2$  (equal priors)

$$\mu_1 = \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \Sigma_1 = \begin{bmatrix} 2 & 0 \\ 0 & 1 \end{bmatrix} \quad \mu_2 = \begin{bmatrix} 0 \\ 2 \end{bmatrix}, \Sigma_2 = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \mu_i)^T \Sigma_i^{-1} (\mathbf{x} - \mu_i) - \frac{1}{2} \ln |\Sigma_i| + \ln P(\omega_i)$$

$$g_1(\mathbf{x}) = -\frac{1}{2} \begin{bmatrix} x & y \end{bmatrix} \begin{bmatrix} 1/2 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} x \\ y \end{bmatrix} - \frac{1}{2} \ln 2$$

$$g_2(\mathbf{x}) = -\frac{1}{2} \begin{bmatrix} x & (y-2) \end{bmatrix} \begin{bmatrix} x \\ (y-2) \end{bmatrix}$$

### Example of decision boundary derivation

$$g_1(\mathbf{x}) = g_2(\mathbf{x})$$

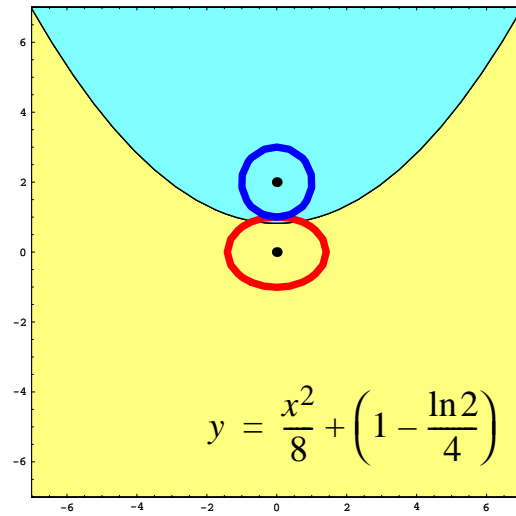
$$-\frac{1}{2} \begin{bmatrix} x & y \end{bmatrix} \begin{bmatrix} 1/2 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} x \\ y \end{bmatrix} - \frac{1}{2} \ln 2 = -\frac{1}{2} \begin{bmatrix} x & (y-2) \end{bmatrix} \begin{bmatrix} x \\ (y-2) \end{bmatrix}$$

$$\begin{bmatrix} x & y \end{bmatrix} \begin{bmatrix} 1/2 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} x \\ y \end{bmatrix} + \ln 2 = \begin{bmatrix} x & (y-2) \end{bmatrix} \begin{bmatrix} x \\ (y-2) \end{bmatrix}$$

$$x^2/2 + y^2 + \ln 2 = x^2 + (y^2 - 4y + 4)$$

$$y = \frac{x^2}{8} + \left(1 - \frac{\ln 2}{4}\right) \text{ (parabola)}$$

## Example of decision boundary derivation



Other derivations: *Mathematica...*