

## Review: Gaussian-based classification

- Conic decision boundaries in feature space
- Limited classification power (*Mathematica*)
  - *Synthetic data examples*
  - *Object classification*

More expressive modeling approach: **mixture modeling**

## Mixture modeling

Basic model:

$$p(\mathbf{x}|\theta) = \sum_{i=1}^k p(\mathbf{x}|\phi_i)P(\omega_i)$$

where,

$$p(\mathbf{x}|\phi_i) = N(\mathbf{x}, \mu_i, \Sigma_i) = \text{Normal density}$$

$$P(\omega_i) = \text{probability (weight) of } i\text{th Normal density}$$

$$\phi_i = \{\mu_i, \Sigma_i\}$$

Given data, what needs to be estimated?

## Mixture modeling: a look ahead

- **No closed form maximum-likelihood solution**
- **Iterative training (Expectation-Maximization)**
- **Random initialization of parameters.**

Some examples before the math (*Mathematica*)

- *Synthetic data examples*
- *Object classification*

## Mixture modeling

*Ok, but...*

*...how does the mixture-model estimation work?  
(Why no closed-form solution?)*

## Simple two-parameter example

Example: 2-Gaussian mixture with known  $\sigma$ , and  $P(\omega_i)$ :

- $P(\omega_1) = 1/3, P(\omega_2) = 2/3$
- $\sigma_1 = \sigma_2 = 1$

Need to estimate  $\{\mu_1, \mu_2\}$  for some data  $\mathbf{X} = \{x_1, \dots, x_n\}$

## Simple two-parameter example

$$p(\mathbf{X}|\mu_1, \mu_2) = \prod_{j=1}^n [p(x_j|\mu_1)P(\omega_1) + p(x_j|\mu_2)P(\omega_2)]$$

$$\ln p(\mathbf{X}|\mu_1, \mu_2) = \sum_{j=1}^n \ln [p(x_j|\mu_1)P(\omega_1) + p(x_j|\mu_2)P(\omega_2)]$$

$$\ln p(\mathbf{X}|\mu_1, \mu_2) = \sum_{j=1}^n \ln \left\{ \exp\left[-\frac{1}{2}(x_j - \mu_1)^2\right]P(\omega_1) + \exp\left[-\frac{1}{2}(x_j - \mu_2)^2\right]P(\omega_2) \right\} - n \ln \sqrt{2\pi}$$

## Simple two-parameter example

*Can this be simplified?*

$$\ln p(\mathbf{X}|\mu_1, \mu_2) = \sum_{j=1}^n \ln \left\{ \exp\left[-\frac{1}{2}(x_j - \mu_1)^2\right]P(\omega_1) + \exp\left[-\frac{1}{2}(x_j - \mu_2)^2\right]P(\omega_2) \right\} - n \ln \sqrt{2\pi}$$

Numeric example (*Mathematica*)

## Maximum-likelihood estimation: mixture models

**Conclusion:** no closed-form solution to maximum-likelihood estimation for mixture models

Could use gradient-based algorithm; however, there is a better answer: **Expectation-Maximization**

## From before: Maximum-likelihood parameter estimation for single Gaussian

For  $d$ -dimensional Gaussians:

$$\mu^* = \frac{1}{n} \sum_{j=1}^n \mathbf{x}_j$$

$$\Sigma^* = \frac{1}{n} \sum_{j=1}^n (\mathbf{x}_j - \mu)(\mathbf{x}_j - \mu)^T$$

Familiar estimates of  $\mu$  and  $\Sigma$  are maximum-likelihood estimates.

## Understanding the EM algorithm

- Parameter estimation of mixture model would be no problem if...?
- Introduce hidden random variables  $y_{ij}$ :

$$y_{ij} = \begin{cases} 1 & \mathbf{x}_j \text{ came from class } \omega_i \\ 0 & \text{otherwise} \end{cases}$$

- Each data point has known and “hidden component.”
- “Complete data” is given by:

$$\mathbf{z}_j = \{\mathbf{x}_j, y_{1j}, y_{2j}, \dots, y_{kj}\}, j \in \{1, 2, \dots, n\}$$

## Expressions for single Normal density

$$\mu_i^* = \frac{1}{n_i} \sum_{j=1}^{n_i} \mathbf{x}_j^{(i)} = \left( \sum_{j=1}^n y_{ij} \mathbf{x}_j \right) / \left( \sum_{j=1}^n y_{ij} \right)$$

$$\begin{aligned} \Sigma_i^* &= \frac{1}{n_i} \sum_{j=1}^{n_i} (\mathbf{x}_j^{(i)} - \mu_i)(\mathbf{x}_j^{(i)} - \mu_i)^T \\ &= \left( \sum_{j=1}^n y_{ij} (\mathbf{x}_j - \mu_i)(\mathbf{x}_j - \mu_i)^T \right) / \left( \sum_{j=1}^n y_{ij} \right) \end{aligned}$$

$$P(\omega_i)^* = \frac{n_i}{n} = \left( \sum_{j=1}^n y_{ij} \right) / n$$

## Very informal statement of EM

However, we don't know the unobserved  $y_{ij}$  variables ...  
EM to the rescue!

**Expectation step:** Calculate the expected value  $E[y_{ij} | \Theta^{(t)}]$  for the hidden variables  $y_{ij}$ , given the current estimate of the parameters at step  $t$  ( $\Theta^{(t)}$ ).

**Maximization step:** Calculate a new maximum-likelihood estimate for the parameters  $(\mu_i^{(t+1)}, \Sigma_i^{(t+1)}, P(\omega_i))$  assuming that the value taken on by each hidden variable  $y_{ij}$  is its expected value  $E[y_{ij} | \Theta^{(t)}]$  (calculated in the Expectation step). Then, iterate.

## Expectation-Maximization algorithm

$$\mu_i^{(t+1)} = \left( \sum_{j=1}^n E[y_{ij} | \Theta^{(t)}] \mathbf{x}_j \right) / \left( \sum_{j=1}^n E[y_{ij} | \Theta^{(t)}] \right)$$

$$\Sigma_i^{(t+1)} = \frac{\sum_{j=1}^n E[y_{ij} | \Theta^{(t)}] (\mathbf{x}_j - \mu_i) (\mathbf{x}_j - \mu_i)^T}{\sum_{j=1}^n E[y_{ij} | \Theta^{(t)}]}$$

$$P(\omega_i)^{(t+1)} = \frac{1}{n} \sum_{j=1}^n E[y_{ij} | \Theta^{(t)}]$$

Let's see how this works...

*(Mathematica)*

**What is  $E[y_{ij} | \Theta^{(t)}]$  ?**

**By definition:**

$$E[y_{ij} | \Theta^{(t)}] = \sum_{x=0}^1 x P(y_{ij} = x | \Theta^{(t)})$$

**So:**

$$E[y_{ij} | \Theta^{(t)}] = 0 \times P(y_{ij} = 0 | \Theta^{(t)}) + 1 \times P(y_{ij} = 1 | \Theta^{(t)})$$

$$E[y_{ij} | \Theta^{(t)}] = P(y_{ij} = 1 | \Theta^{(t)})$$

$$E[y_{ij} | \Theta^{(t)}] = P(\omega_i | \mathbf{x}_j, \Theta^{(t)})$$

**So what is  $P(\omega_i | \mathbf{x}_j, \Theta^{(t)})$  ?**

$$P(\omega_i | \mathbf{x}_j, \Theta^{(t)}) = \frac{p(\mathbf{x}_j | \omega_i, \Theta^{(t)}) P(\omega_i | \Theta^{(t)})}{p(\mathbf{x}_j | \Theta^{(t)})} \quad \text{(Bayes Rule)}$$

**where:**

$$p(\mathbf{x}_j | \omega_i, \Theta^{(t)}) = N[\mathbf{x}_j, \mu_i^{(t)}, \Sigma_i^{(t)}], \quad P(\omega_i | \Theta^{(t)}) = P(\omega_i)^{(t)}$$

$$\begin{aligned} p(\mathbf{x}_j | \Theta^{(t)}) &= \sum_{i=1}^k p(\mathbf{x}_j | \omega_i, \Theta^{(t)}) P(\omega_i | \Theta^{(t)}) \\ &= \sum_{i=1}^k N[\mathbf{x}_j, \mu_i^{(t)}, \Sigma_i^{(t)}] P(\omega_i)^{(t)} \end{aligned}$$

## Expectation-Maximization algorithm

$$\mu_i^{(t+1)} = \left( \sum_{j=1}^n E[y_{ij} | \Theta^{(t)}] \mathbf{x}_j \right) / \left( \sum_{j=1}^n E[y_{ij} | \Theta^{(t)}] \right)$$

$$\Sigma_i^{(t+1)} = \frac{\sum_{j=1}^n E[y_{ij} | \Theta^{(t)}] (\mathbf{x}_j - \mu_i) (\mathbf{x}_j - \mu_i)^T}{\sum_{j=1}^n E[y_{ij} | \Theta^{(t)}]}$$

$$P(\omega_i)^{(t+1)} = \frac{1}{n} \sum_{j=1}^n E[y_{ij} | \Theta^{(t)}]$$

## Expected value of hidden variables

$$E[y_{ij} | \Theta^{(t)}] = \frac{p(\mathbf{x}_j | \omega_i, \Theta^{(t)}) P(\omega_i | \Theta^{(t)})}{p(\mathbf{x}_j | \Theta^{(t)})}$$

where:

$$p(\mathbf{x}_j | \omega_i, \Theta^{(t)}) = N[\mathbf{x}_j, \mu_i^{(t)}, \Sigma_i^{(t)}], \quad P(\omega_i | \Theta^{(t)}) = P(\omega_i)^{(t)}$$

$$\begin{aligned} p(\mathbf{x}_j | \Theta^{(t)}) &= \sum_{i=1}^k p(\mathbf{x}_j | \omega_i, \Theta^{(t)}) P(\omega_i | \Theta^{(t)}) \\ &= \sum_{i=1}^k N[\mathbf{x}_j, \mu_i^{(t)}, \Sigma_i^{(t)}] P(\omega_i)^{(t)} \end{aligned}$$

## Informal definition of EM

1. Choose an initial estimate for  $\Theta$ .
2. **(E)xpectation step:** Calculate the expected value  $E[y_{ij} | \Theta]$  for the hidden variables  $y_{ij}$ , given the current estimate of the parameters ( $\Theta$ ).
3. **(M)aximization step:** Calculate a new maximum-likelihood estimate for the parameters ( $\Theta$ ) assuming that the value taken on by each hidden variable  $y_{ij}$  is its expected value  $E[y_{ij} | \Theta]$  (calculated in the Expectation step).
4. Iterate **Expectation** and **Maximization** steps.

## Formal problem statement for EM algorithm

**Definitions:**

- Let  $\mathbf{X} = \{\mathbf{x}_j\}$ ,  $j \in \{1, 2, \dots, n\}$  denote  $n$  *observed* (incomplete) data vectors.
- Let  $\mathbf{Y} = \{\mathbf{y}_j\}$ ,  $j \in \{1, 2, \dots, n\}$  denote the corresponding  $n$  *unobserved (hidden)* data vectors.
- Let  $\mathbf{Z} = \mathbf{X} \cup \mathbf{Y}$  denote the *full (or complete)* data where,

$$\mathbf{z}_j = \{\mathbf{x}_j, \mathbf{y}_j\}$$

**Goal:** Maximize  $L(\mathbf{X}, \Theta)$  with respect to  $\Theta$  such that,

$$L(\mathbf{X}, \Theta) \leq L(\mathbf{X}, \Theta^*), \quad \forall \Theta.$$

## The EM algorithm (huh?)

1. Choose an initial estimate for  $\Theta$ .

2. (E)xpectation step: Compute  $Q(\Theta, \bar{\Theta})$ ,

$$Q(\Theta, \bar{\Theta}) = E[L(\mathbf{Z}|\bar{\Theta})|\mathbf{X}, \Theta], \text{ where,}$$

$$L(\mathbf{Z}|\bar{\Theta}) = \sum_{j=1}^n \ln p(\mathbf{x}_j, \mathbf{y}_j|\bar{\Theta}).$$

3. (M)aximization step: Replace the current estimate  $\Theta$  with the new estimate  $\bar{\Theta}$  where,

$$\bar{\Theta} = \underset{\Theta}{\operatorname{argmax}} Q(\Theta, \bar{\Theta})$$

4. Iterate steps 2 and 3 until convergence.

## Some discussion on the EM algorithm

- Huh?
- Important property: maximizing  $Q(\Theta, \bar{\Theta})$  maximizes  $L(\mathbf{X}, \bar{\Theta})$
- Proof of result for mixture modeling is non-trivial.

## Some key points

Question: How are  $Q(\Theta, \bar{\Theta})$  and  $L(\mathbf{X}, \bar{\Theta})$  related?

$$L(\mathbf{X}|\Theta) \equiv \sum_{j=1}^n \ln p(\mathbf{x}_j|\Theta)$$

Let's express  $p(\mathbf{x}|\Theta)$  in terms of  $\mathbf{x}$  and  $\mathbf{y}$ :

$$p(\mathbf{y}|\mathbf{x}) \equiv \frac{p(\mathbf{x}, \mathbf{y})}{p(\mathbf{x})} \Rightarrow p(\mathbf{y}|\mathbf{x}, \Theta) = \frac{p(\mathbf{x}, \mathbf{y}|\Theta)}{p(\mathbf{x}|\Theta)}$$

$$p(\mathbf{x}|\Theta) = \frac{p(\mathbf{x}, \mathbf{y}|\Theta)}{p(\mathbf{y}|\mathbf{x}, \Theta)}$$

$$\ln p(\mathbf{x}|\Theta) = \ln p(\mathbf{x}, \mathbf{y}|\Theta) - \ln p(\mathbf{y}|\mathbf{x}, \Theta)$$

## How are $Q(\Theta, \bar{\Theta})$ and $L(\mathbf{X}, \bar{\Theta})$ related?

$$\ln p(\mathbf{x}|\Theta) = \ln p(\mathbf{x}, \mathbf{y}|\Theta) - \ln p(\mathbf{y}|\mathbf{x}, \Theta)$$

$$\ln p(\mathbf{x}|\Theta) = \ln p(\mathbf{z}|\Theta) - \ln p(\mathbf{y}|\mathbf{x}, \Theta)$$

$$\sum_{j=1}^n \ln p(\mathbf{x}_j|\Theta) = \sum_{j=1}^n \ln p(\mathbf{z}_j|\Theta) - \sum_{j=1}^n \ln p(\mathbf{y}_j|\mathbf{x}_j, \Theta)$$

$$L(\mathbf{X}|\Theta) = L(\mathbf{Z}|\Theta) - L(\mathbf{Y}|\mathbf{X}, \Theta)$$

Change  $\Theta$  to  $\bar{\Theta}$  and apply  $E[\cdot|\mathbf{X}, \Theta]$ :

$$L(\mathbf{X}|\bar{\Theta}) = L(\mathbf{Z}|\bar{\Theta}) - L(\mathbf{Y}|\mathbf{X}, \bar{\Theta})$$

$$E[L(\mathbf{X}|\bar{\Theta})|\mathbf{X}, \Theta] = E[L(\mathbf{Z}|\bar{\Theta})|\mathbf{X}, \Theta] - E[L(\mathbf{Y}|\mathbf{X}, \bar{\Theta})|\mathbf{X}, \Theta]$$

### How are $Q(\Theta, \bar{\Theta})$ and $L(\mathbf{X}, \bar{\Theta})$ related?

$$E[L(\mathbf{X}|\bar{\Theta})|\mathbf{X}, \Theta] = E[L(\mathbf{Z}|\bar{\Theta})|\mathbf{X}, \Theta] - E[L(\mathbf{Y}|\mathbf{X}, \bar{\Theta})|\mathbf{X}, \Theta]$$

**Note:**

$$Q(\Theta, \bar{\Theta}) \equiv E[L(\mathbf{Z}|\bar{\Theta})|\mathbf{X}, \Theta]$$

**so that:**

$$E[L(\mathbf{X}|\bar{\Theta})|\mathbf{X}, \Theta] = Q(\Theta, \bar{\Theta}) - E[L(\mathbf{Y}|\mathbf{X}, \bar{\Theta})|\mathbf{X}, \Theta]$$

### How are $Q(\Theta, \bar{\Theta})$ and $L(\mathbf{X}, \bar{\Theta})$ related?

$$E[L(\mathbf{X}|\bar{\Theta})|\mathbf{X}, \Theta] = Q(\Theta, \bar{\Theta}) - E[L(\mathbf{Y}|\mathbf{X}, \bar{\Theta})|\mathbf{X}, \Theta]$$

**What is  $E[L(\mathbf{X}|\bar{\Theta})|\mathbf{X}, \Theta]$ ?**

- $L(\mathbf{X}|\bar{\Theta})$  is independent of  $\Theta$  (constant for given  $\mathbf{X}, \bar{\Theta}$ )
- $E[L(\mathbf{X}|\bar{\Theta})|\mathbf{X}, \Theta] = L(\mathbf{X}|\bar{\Theta})$

**So:**

$$L(\mathbf{X}|\bar{\Theta}) = Q(\Theta, \bar{\Theta}) - E[L(\mathbf{Y}|\mathbf{X}, \bar{\Theta})|\mathbf{X}, \Theta]$$

### How are $Q(\Theta, \bar{\Theta})$ and $L(\mathbf{X}, \bar{\Theta})$ related?

$$L(\mathbf{X}|\bar{\Theta}) = Q(\Theta, \bar{\Theta}) - E[L(\mathbf{Y}|\mathbf{X}, \bar{\Theta})|\mathbf{X}, \Theta]$$

**Define:**

$$H(\Theta, \bar{\Theta}) \equiv E[L(\mathbf{Y}|\mathbf{X}, \bar{\Theta})|\mathbf{X}, \Theta]$$

$$L(\mathbf{X}|\bar{\Theta}) = Q(\Theta, \bar{\Theta}) - H(\Theta, \bar{\Theta})$$

**Can be shown that (see notes):**

$$H(\Theta, \bar{\Theta}) \leq H(\Theta, \Theta)$$

**Therefore:**

$$Q(\Theta, \bar{\Theta}) \geq Q(\Theta, \Theta) \text{ implies } L(\mathbf{X}|\bar{\Theta}) \geq L(\mathbf{X}|\Theta).$$

### Expectation step: expression for $Q(\Theta, \bar{\Theta})$

**By definition:**

$$Q(\Theta, \bar{\Theta}) \equiv E[L(\mathbf{Z}|\bar{\Theta})|\mathbf{X}, \Theta]$$

**What does  $E[\bullet|\mathbf{X}, \Theta]$  mean?**

**Simpler example:**

- $f(\mathbf{x}, \mathbf{y})$ ,  $\mathbf{x}$  = known constant,  $\mathbf{y}$  = random variable(s)
- $E[f(\mathbf{x}, \mathbf{y})] \equiv \sum_{\mathbf{y}} f(\mathbf{x}, \mathbf{y})p(\mathbf{y})$  (expected value)
- $E[f(\mathbf{x}, \mathbf{y})|\mathbf{x}, \Theta] \equiv \sum_{\mathbf{y}} f(\mathbf{x}, \mathbf{y})p(\mathbf{y}|\mathbf{x}, \Theta)$  (conditional expected value)

## Expression for $Q(\Theta, \bar{\Theta})$

By definition:

$$Q(\Theta, \bar{\Theta}) \equiv E[L(\mathbf{Z}|\bar{\Theta})|\mathbf{X}, \Theta]$$

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \left( \sum_{\mathbf{y}_j} \ln p(\mathbf{x}_j, \mathbf{y}_j|\bar{\Theta}) p(\mathbf{y}_j|\mathbf{x}_j, \Theta) \right).$$

$$p(\mathbf{y}_j|\mathbf{x}_j, \Theta) = \frac{p(\mathbf{x}_j, \mathbf{y}_j|\Theta)}{p(\mathbf{x}_j|\Theta)}$$

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \sum_{\mathbf{y}_j} \ln p(\mathbf{x}_j, \mathbf{y}_j|\bar{\Theta}) \frac{p(\mathbf{x}_j, \mathbf{y}_j|\Theta)}{p(\mathbf{x}_j|\Theta)}$$

## Expression for $Q(\Theta, \bar{\Theta})$

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \sum_{\mathbf{y}_j} \ln p(\mathbf{x}_j, \mathbf{y}_j|\bar{\Theta}) \frac{p(\mathbf{x}_j, \mathbf{y}_j|\Theta)}{p(\mathbf{x}_j|\Theta)}$$

For mixture model:

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \sum_{i=1}^k \ln [p(\mathbf{x}_j|\bar{\phi}_i) \overline{P(\omega_i)}] \frac{p(\mathbf{x}_j|\phi_i) P(\omega_i)}{p(\mathbf{x}_j|\Theta)}$$

(why?)

## Expression for $Q(\Theta, \bar{\Theta})$ (mixture model)

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \sum_{i=1}^k \ln [p(\mathbf{x}_j|\bar{\phi}_i) \overline{P(\omega_i)}] \frac{p(\mathbf{x}_j|\phi_i) P(\omega_i)}{p(\mathbf{x}_j|\Theta)}$$

Note:

$$\frac{p(\mathbf{x}_j|\phi_i) P(\omega_i)}{p(\mathbf{x}_j|\Theta)} = P(\omega_i|\mathbf{x}_j, \Theta)$$

so that:

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \sum_{i=1}^k P(\omega_i|\mathbf{x}_j, \Theta) \ln [p(\mathbf{x}_j|\bar{\phi}_i) \overline{P(\omega_i)}]$$

## Expectation step (mixture model)

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \sum_{i=1}^k P(\omega_i|\mathbf{x}_j, \Theta) \ln [p(\mathbf{x}_j|\bar{\phi}_i) \overline{P(\omega_i)}]$$

Final form of  $Q$ -function:

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \sum_{i=1}^k P(\omega_i|\mathbf{x}_j, \Theta) \ln \overline{P(\omega_i)} + \sum_{j=1}^n \sum_{i=1}^k P(\omega_i|\mathbf{x}_j, \Theta) \ln p(\mathbf{x}_j|\bar{\phi}_i)$$



## Maximization step (mixture model)

To maximize w/respect to  $\bar{\Theta}$  :

$$\nabla_{\bar{\Theta}} Q(\Theta, \bar{\Theta}) = 0$$

Two terms:

$$\sum_{j=1}^n \sum_{i=1}^k P(\omega_i | \mathbf{x}_j, \Theta) \ln \overline{P(\omega_i)}$$

$\Rightarrow$  constrained Lagrangian optimization:

$$\sum_{i=1}^k \overline{P(\omega_i)} = 1$$

## Maximization step (mixture model)

Second term:

$$\sum_{j=1}^n \sum_{l=1}^k P(\omega_l | \mathbf{x}_j, \Theta) \ln p(\mathbf{x}_j | \bar{\Phi}_l)$$

$\Rightarrow$  set gradient w/respect to  $\bar{\Phi}_i$  to zero and solve:

$$\nabla_{\bar{\Phi}_i} \sum_{j=1}^n \sum_{l=1}^k P(\omega_l | \mathbf{x}_j, \Theta) \ln p(\mathbf{x}_j | \bar{\Phi}_l) = 0$$

$$\sum_{j=1}^n P(\omega_i | \mathbf{x}_j, \Theta) \frac{\nabla_{\bar{\Phi}_i} p(\mathbf{x}_j | \bar{\Phi}_i)}{p(\mathbf{x}_j | \bar{\Phi}_i)} = 0$$

## Maximization step (mixture model)

$$\sum_{j=1}^n P(\omega_i | \mathbf{x}_j, \Theta) \frac{\nabla_{\bar{\Phi}_i} p(\mathbf{x}_j | \bar{\Phi}_i)}{p(\mathbf{x}_j | \bar{\Phi}_i)} = 0$$

For exponential  $p(\mathbf{x}_j | \bar{\Phi}_i)$  :

$$\nabla_{\bar{\Phi}_i} p(\mathbf{x}_j | \bar{\Phi}_i) \Rightarrow f(\mathbf{x}_j, \bar{\Phi}_i) p(\mathbf{x}_j | \bar{\Phi}_i)$$

so that:

$$\sum_{j=1}^n P(\omega_i | \mathbf{x}_j, \Theta) f(\mathbf{x}_j, \bar{\Phi}_i) = 0$$

This property typically allows us to solve for  $\bar{\Phi}_i$  ...

## Final solution of EM update equations for mixture models

*many equations later...*

$$\overline{P(\omega_i)} = \frac{1}{n} \sum_{j=1}^n P(\omega_i | \mathbf{x}_j, \Theta)$$

$$\bar{\mu}_i = \left( \sum_{j=1}^n P(\omega_i | \mathbf{x}_j, \Theta) \mathbf{x}_j \right) / \left( \sum_{j=1}^n P(\omega_i | \mathbf{x}_j, \Theta) \right)$$

$$\bar{\Sigma}_i = \sum_{j=1}^n P(\omega_i | \mathbf{x}_j, \Theta) (\mathbf{x}_j - \bar{\mu}_i)(\mathbf{x}_j - \bar{\mu}_i)^T / \sum_{j=1}^n P(\omega_i | \mathbf{x}_j, \Theta)$$

## EM Summary

- **General convergence proof**

$$Q(\Theta, \bar{\Theta}) \geq Q(\Theta, \Theta) \text{ implies } L(\mathbf{X}|\bar{\Theta}) \geq L(\mathbf{X}|\Theta).$$

- **Expressions for  $Q(\Theta, \bar{\Theta})$  (mixture models):**

$$Q(\Theta, \bar{\Theta}) = \sum_{j=1}^n \sum_{i=1}^k P(\omega_i | \mathbf{x}_j, \Theta) \ln \overline{P(\omega_i)} + \sum_{j=1}^n \sum_{i=1}^k P(\omega_i | \mathbf{x}_j, \Theta) \ln p(\mathbf{x}_j | \bar{\Phi}_i)$$

## EM concluding thoughts...

### Powerful algorithm for training complex statistical models:

- Mixture models (Gaussian, non-Gaussian example in new notes)
- Vector quantization (next)
- Hidden Markov models (temporal mixture models)
- Bayesian belief networks, Markov tree models, etc...

### Applications:

- Computer vision (object detection/recognition)
- Speech recognition
- Gesture recognition, human performance analysis, etc....